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Geographical heterogeneity and influenza infection within households

Day-Yu Chao^{1*}, Kuang-Fu Cheng², Ying-Hen Hsieh^{3,4}, Tsai-Chung Li^{3,4}, Trong-Neng Wu^{3,4} and Chiu-Ying Chen⁴

Abstract

Background: Although it has been suggested that schoolchildren vaccination reduces influenza morbidity and mortality in the community, it is unknown whether geographical heterogeneity would affect vaccine effectiveness.

Methods: A 3-year prospective, non-randomized sero-epidemiological study was conducted during 2008–2011 by recruiting schoolchildren from both urban and rural areas. Respective totals of 124, 206, and 176 households were recruited and their household contacts were followed. Serum samples were collected pre-vaccination, one-month post-vaccination and post-season from children and household contacts for hemagglutination inhibition (HI) assay. A multivariate logistic model implemented with generalized estimation equations (GEE) was fitted with morbidity or a four-fold increase in HI titer of the household contacts for two consecutive sera as the dependent variable; with geographical location, vaccination status of each household and previous vaccination history as predictor variables.

Results: Although our results show no significant reduction in the proportion of infection or clinical morbidity among household contacts, a higher risk of infection, indicated by odds ratio > 1, was consistently observed among household children contacts from the un-vaccinated households after adjusting for confounding variables. Interestingly, a statistically significant lower risk of infection was observed among household adult contacts from rural area when compared to those from urban area (OR = 0.89; 95% CI: 0.82-0.97 for Year 2 and OR = 0.85; 95% CI: 0.75-0.96 for Year 3).

Conclusions: A significant difference in the risk of influenza infection among household adults due to geographical heterogeneity, independent of schoolchildren vaccination status, was revealed in this study. Its impact on vaccine effectiveness requires further study.

Keywords: Influenza, Trivalent Inactivated Vaccine (TIV), Children, Household contacts, Geographical heterogeneity

Background

Influenza is a major cause of morbidity and mortality, resulting in an estimated 3–5 million cases of severe influenza illness annually [1]. Although older adults have the highest influenza-related mortality, children who have contracted influenza infection experience substantial morbidity, resulting in absence from school, extra working days for parents and increased health care costs from purchasing antibiotics [2-4]. Additionally, children attending day-care centers and elementary schools have long been identified as the major causes of influenza virus transmission in the community since they can shed greater amounts of virus for longer periods of time [5-8]. In fact,

the best predictor for influenza occurring in a household is the presence of children [6]. Focusing efforts for influenza vaccination on school-aged children may therefore be an effective and practical method for reducing the burden of influenza in the community. Since 2008, the Advisory Committee on Immunization Practices (ACIP) of the Centers for Disease Control and Prevention (CDC) in the U.S.A. expanded universal influenza vaccination recommendations to all children aged 0.5-18 years [9].

A systematic review suggested that although evidence exists that vaccinating healthy children has the potential of reducing the effect of influenza transmission within households and community, further data are needed because of limitations in study design, varied vaccination policy implemented by different countries and prior experience of receiving vaccination which make the benefits

¹Graduate Institute of Microbiology and Public Health, College of Veterinary Medicine, National Chung-Hsing University, Taichung, Taiwan Full list of author information is available at the end of the article



^{*} Correspondence: dychao@nchu.edu.tw

from vaccination difficult to quantify [10-14]. Furthermore, recent studies from 2009 pandemic H1N1 (pH1N1) outbreaks suggest that transmissions of influenza virus were spatially heterogeneous [15]. So far, it is unknown whether spatially heterogeneous transmission of influenza virus affects the infection rate and vaccine effectiveness within households.

Since 2007, the government in Taiwan has implemented a free vaccination program through which all schoolchildren in grades 1-4 receive a free annual trivalent influenza vaccination (TIV), which later expanded to grades 1-6 after 2010: a single dose from commercially-available TIV containing 0.5-mL of 15ug HA of the H1N1, H3N2 and B antigen would be administered through a school-based delivery program. From 2008-2011, we recruited elementary schoolchildren from both urban and rural areas during three consecutive influenza seasons and followed up three different cohorts of their household contacts. The current non-randomized study offered a unique opportunity to evaluate (1) the effect of vaccinating school-aged children on reducing virus transmission and influenza related morbidity among their household contacts, and (2) the influence of geographical location on acquiring influenza infection among the household contacts.

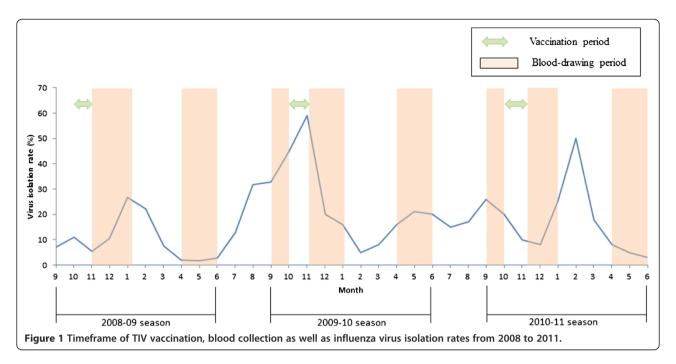
Methods

Subject enrollment and serological specimens

During three consecutive influenza seasons from 2008 to 2011, a prospective non-randomized sero-epidemiological study was conducted by the influenza research group at National Chung-Hsing University (NCHU) and China Medical University (CMU) to investigate household transmission and vaccine effectiveness. Students from two urban (Taichung City) elementary schools and four rural (Nantou County) schools in central Taiwan were recruited for this three-year study starting in the fall of 2008. Taichung city is the largest urban city in central Taiwan with a population of more than 1 million and has a highly developed socioeconomic structure. The two schools selected were located respectively in the North and Central districts of the city, with approximately 140 thousand total residents. The nearby Nantou County, which has a total land size approximately 25 times larger than that of Taipei City, is the second largest county and the only landlocked county on the island of Taiwan with a population of over 500 thousand, and is comparatively less developed socioeconomically. The four schools in Nantou County were selected purposely from four different administrative districts in the county; namely, Nantou City and Tsaotun Township each with around 100 thousand residents, and the rural townships of, Mingjian, and Guoshing with 40 and 20 thousand residents, respectively. Since the study protocol involved intensive visiting and blood-drawing, all children from the selected schools of grades 1-4 during the first and second study year and grades 1-6 during the third study year were given a detailed description of the study protocol and a returned slip was included to acknowledge if they were willing to participate in the study voluntarily. Each successive year, some families dropped out and additional volunteer subjects were recruited. Family members of the school-aged children were also recruited to join the study to further understand vaccine effectiveness in preventing household transmissions of seasonal influenza viruses. The serum samples including pre-vaccination (collected between September and October), one month post-vaccination (collected between November and January) and postseason (collected between April and June) were taken from children participants by trained nurses, except during the 2008-2009 season when only sera from postvaccination and post-season periods were taken as shown in Figure 1. Since TIV was not free for household adult contacts unless they were aged greater than 65 years and the proportion of TIV vaccination was less than 0.1%, sera were taken from adult contacts only during pre- and postvaccination during the three consecutive seasons. All subjects gave informed consent and the study was approved by the Medical Ethics Committee of China Medical University (DMR96-IRB-216). The study protocol based on laboratory data was established to assess the serological responses by including the paired sera drawn from two consecutive time periods, in order to define immune response before and after vaccination or before vaccination and post-season among household contacts. The obtained sera were evaluated for antibody titers by measuring a hemagglutination inhibition (HI) assay following the standard protocol by the World Health Organization (WHO) using the contemporary vaccine and wild-type strains of influenza viruses as described in the laboratory methods section below.

Data collection

All study subjects including children participants and household contacts were assessed for signs and symptoms of influenza throughout the follow-up period, defined by the start date (>1 laboratory-confirmed influenza cases in 2 consecutive weeks from sentinel sites) and termination date (no laboratory-confirmed influenza cases for 2 consecutive weeks from sentinel sites) based on the influenza surveillance network coordinated by Taiwan-CDC which covers approximately 75% of the basic administrative units of Taiwan (cities or districts) and routinely collects clinical specimens [16,17]. The average influenza virus isolation rates per month during three consecutive seasons were shown in Figure 1. During each influenza season, two questionnaires were also administered by trained interviewers regarding basic



demographic and social contact information and whether a seasonal influenza vaccination had been received in the past and present years. Information regarding comorbidities including cardiovascular disease, hypertension, or diabetes mellitus was also taken from the adults in the family. Bi-weekly clinical symptom reviews were carried out using a standardized questionnaire via a telephone interview between October and June. Participants were asked to report any newly experienced febrile respiratory symptoms, including fevers ($\geq 38^{\circ}$ C), sore throats, cough, nasal congestion, headache, sinus problems, muscle aches, fatigue, ear ache or infection or chills during the influenza season. The overall completion rate for the questionnaires was more than 90%.

Laboratory methods

Antibody titers were measured using HI assay which adhered to the WHO protocol [18,19]. In brief, serum samples were pre-treated with receptor destroying enzyme (RDE, Deka Seriken Co Ltd, Tokyo, Japan) in 1:4 ratio at 37°C for 16 hours, followed by another 30 minutes at 56°C and then an equal volume of 1.6% trisodium citrate was added for enzyme inactivation. The different strains of influenza viruses used in this study were first prepared from the culture supernatants of infected Madin-Darby canine kidney (MDCK) cells. Twenty-five microliters (µl) (4 hemagglutination units, HA) of influenza virus were incubated at room temperature for one hour with an equal volume of RDE-treated serum in a V-shape 96well microtiter plate. After incubation, 25 µl of 1% (vol/ vol) chicken red blood cells was added to each well. Hemagglutination inhibition was read after 30 minutes. To evaluate asymptomatic infection, influenza viruses from the contemporary vaccine and wild-type strains which represented more than 80% of circulating viruses during the three consecutive influenza seasons were used. Details are listed in Table 1. For the HI assay, serum samples were tested with an initial dilution of 1:10 and a final dilution of 1:10,240 and the titers were expressed as the reciprocal of the highest dilution of serum where hemagglutination was prevented. Samples with dilution higher than 10,240 where hemagglutination was prevented were repeated to obtain the final HI titers. Samples that were negative by HI were assigned a titer of 1:5 for computational purposes in obtaining a geometric mean titer (GMT) or seroconversion rate. We defined seroconversion as a four-fold or greater rise of HI antibody titers to influenza A (H1N1), A (H3N2) and B to assess serological evidence of viral infection among household contacts. In this study, we excluded household contacts who had TIV vaccination during the influenza seasons in the analysis to avoid false classification. This allows us to use a four-fold increase in the sera between pre- and post-season as a surrogate marker for viral infection, and compare the likelihood of infection among household contacts between households of different vaccination statuses as later described in the statistical analysis section.

Statistical analysis

Vaccination status of each household was categorized as 'complete', 'partial' or 'non-vaccination' based on the vaccination status of its occupying school-aged child/children (named as participants). If all participants in the

Table 1 The contemporary vaccine and wild-type strains circulated during the three consecutive influenza seasons used in this study

Strain	2008-2009(1st Year)	2009-2010(2nd Year)	2010-2011(3rd Year)
Vaccine strain			
H1N1	A/Brisbane/59/2007	A/Brisbane/59/2007	
pdmH1N12009 ²		A/California/07/2009	A/California/07/2009
H3N2	A/Brisbane/10/2007	A/Brisbane/10/2007	A/Perth/16/2009
flu B	B/Florida/4/2006(Y) ¹	B/Brisbane/60/2008(V) ¹	B/Brisbane/60/2008(V)
Wildtype strain			
H1N1	A/Taiwan/606/2008		
pdmH1N12009 ²		A/California/07/2009	A/Taiwan/5506/2010
			A/Taiwan/5520/2011
H3N2	A/Taiwan/736/2008	A/Taiwan/480/2008	A/Taiwan/3869/2010
	A/Taiwan/480/2008	A/Taiwan/736/2008	A/Taiwan/3814/2011
		A/Taiwan/3982/2009	
flu B	B/Taiwan/29/2008(Y)	B/Taiwan/5908/2009(V)	B/Taiwan/3591/2010(V)
			B/Taiwan/5806/2011(Y)

¹Y refer to Yamagada lineage and V refer to Victoria lineage.

household received or did not receive TIV during the season, the household was classified as 'complete' or 'unvaccination', respectively. If only some of the participants from the household received TIV during the season, the household was classified as 'partial'. To estimate vaccine effectiveness in preventing respiratory illness or infection among household contacts, a logistic model was fitted with the illness status or four-fold increase of HI titer during two consecutive sera of the household contacts as the dependent variable and the vaccination status of each household as the predictor variable. Generalized estimation equations (GEE) method was used to account for correlation among members of the same household. We tested the hypothesis that there was no difference in the rate of respiratory illnesses or related infection among the household contacts of completely vaccinated, partially vaccinated and un-vaccinated households. All statistical analyses were conducted using SAS (release 9.2, SAS Institute) software. A p-value <0.05 was considered statistically significant.

Results

Subject recruitment and demographics

Figure 2 describes the numbers of participants and household contacts recruited in this study during the three consecutive years and the final numbers entered for statistical analysis. Except for the first year, there was more than one school-aged child included in the study from each household.

For households considered in the final analysis, we stratified them into 'complete', 'partial' and 'un-vaccinated' according to the immunization status of the participants

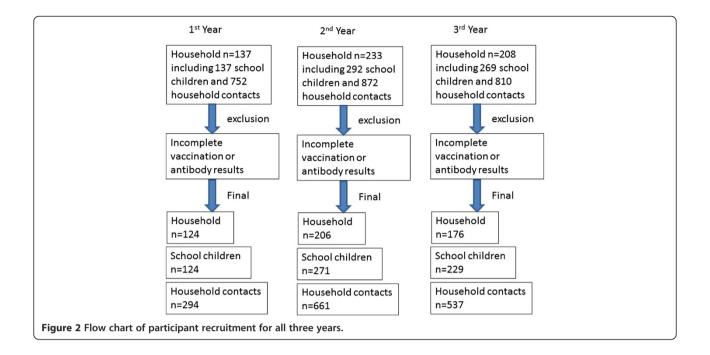
from each household as previously described. Table 2 shows the basic demographics of participants and household contacts from three consecutive seasons. The household sizes were about the same across all three consecutive seasons - the partially-immunized households were slightly larger but not significantly different from the un-vaccinated or completely-vaccinated households. A higher proportion of recruited participants from the urban area (Taichung) did not receive TIV (TIV(–) families) except in Year 2 (51% from TIV(+) vs 30% from TIV(–) and 43% from partial TIV families), probably due to the occurrence of pandemic H1N1 (pH1N1) in 2009.

For children participants from each household, no gender difference was found during all three years. However, it was found that during Years 1 and 3, children participants who received TIV vaccination were slightly younger compared to those who did not. For household contacts, a significantly higher proportion of young contacts (age less than 18 years) was found in unvaccinated (23%) and partially-vaccinated (25%) groups than from the completely-vaccinated group (11%) during the second influenza season (p < 0.05). A significantly higher proportion of household contacts who received TIV during the previous season was also found in 'complete' and 'partial' vaccination households than from 'un-vaccination' households during the first and second year of study (p < 0.05) (Table 2).

Immune response and vaccine efficacy among schoolchildren participants

Table 3 shows the antibody titers from the schoolchildren participants pre- and post-vaccination during three

²pdmH1N12009: pandemic H1N1 in 2009.



consecutive influenza seasons. A significantly higher antibody titer of post-vaccination sera in participants who received TIV vaccination, compared to the titer in participants without receiving TIV, was observed during Year 1 and 3. Meanwhile, significantly higher immune responses, including GMT and proportion of HI equal to or greater than 40, were observed among participants receiving TIV than those without (Table 3). Higher proportions of participants with 4-fold increase of antibody titer were also observed among the group receiving TIV than the group without (51% vs 16% for H1N1 and 16% vs 0% for B antigens). No difference in antibody titers was observed between TIV(+) and TIV(-) participants in the post-season sera for all three study years.

Since pH1N1 circulated before and during the second influenza year, TIV vaccine was delivered one month ahead of pH1N1 vaccination. We further divided the participants from Year 2 with or without receiving TIV into four groups based on their status of receiving pH1N1 vaccination. As shown in Table 4, a significantly higher proportion of participants having HI titer greater than or equal to 40 in the groups receiving the pH1N1 vaccine was observed than in the group without receiving the pH1N1 vaccine, particularly among the groups receiving the pH1N1 but not TIV vaccine (p < 0.05). Similar results were observed when comparing GMT titers between two groups. Interestingly, higher antibody titers against seasonal influenza strains including H1, H3 and B were observed in children who received the pH1N1 vaccine than in children without the pH1N1 vaccine regardless of their TIV history (Table 4).

Effectiveness of vaccination of schoolchildren in reducing illness or serologically proven infection among household contacts

Previous studies suggested that schoolchildren vaccinated with TIV could reduce household and community transmission of influenza virus [10-14]. Our results did not detect significant reduction in the proportion of infection or clinical morbidity among contacts of the households with complete TIV vaccination compared to those with partial or no vaccination (Table 5). Also, we further divided the household based on the status of receiving TIV or pH1N1 vaccines during the 2009-2010 influenza season due to the pH1N1 epidemic. Because of the low numbers of the households did not receive the pH1N1 vaccine, infection rates of contacts from Year 2 were removed. Higher infection rates (38% of type H3 and 84% of any type) were observed among household adult contacts from families receiving both TIV and pH1N1 vaccines than from families not receiving TIV (26% of type H3 and 76% of any type) with statistical significance (p < 0.05) (Table 5). Similar results were observed among household children contacts.

To further investigate if the proportion of serological infection or clinical morbidity of household contacts was associated with other confounding variables, multivariate analysis using GEE was implemented to adjust for variables including underlying diseases, household size, residential location, TIV during the previous year, the status of household receiving TIV. The results were shown in Table 6. Gender of household contacts was not included due to the lack of difference shown

Table 2 Demographic characteristics of trivalent influenza vaccine (TIV) completed, partial and non-vaccinated households during the three consecutive years of influenza seasons (2008–2011)

	1st Year		2nd Year				3rd Year			_	
	TIV(-)	TIV(+)	р	TIV(-)	Partial TIV	TIV(+)	р	TIV(-)	Partial TIV	TIV(+)	р
No of household											
Household size#	3.28	3.45		4.2	7.46	3.84		3.76	7.50	4.12	
	(3.00,3.58)	(3.18,3.72)		(3.83,4.64)	(6.51,8.40)	(3.35,4.34)		(3.15,4.36)	(6.06,8.94)	(3.70,4.54)	
Location of residence											
Taichung	28(0.47)	22(0.38)		34(0.30)	15(0.43)	29(0.51)	*	20(0.49)	9(0.45)	35(0.30)	*
Nantou	31(0.53)	43(0.62)		80(0.70)	20(0.57)	28(0.49)		21(0.51)	11(0.55)	80(0.70)	
Participants											
Gender											
Male	31(0.53)	34(0.52)		71(0.52)	33(0.47)	31(0.48)		19(0.41)	21(0.49)	75(0.54)	
Female	28(0.47)	31(0.48)		66(0.48)	37(0.53)	33(0.52)		27(0.59)	22(0.51)	65(0.46)	
Age (average years)#	9.24	8.00	*	9.39	9.20	9.11		10.80	9.65	9.66	*
	(8.83,9.64)	(7.64,8.36)		(9.10,9.68)	(8.75,9.65)	(8.76,9.45)		(10.24,11.37)	(9.04,10.26)	(9.39,9.94)	
Household contacts											
Gender											
Male	60(0.44)	63(0.40)		152(0.44)	78(0.44)	55(0.40)		40(0.37)	45(0.42)	142(0.44)	
Female	75(0.56)	96(0.60)		192(0.56)	100(0.56)	84(0.60)		68(0.63)	61(0.58)	181(0.56)	
Age											
<18	37(0.27)	43(0.27)		80(0.23)	45(0.25)	15(0.11)	*	31(0.29)	32(0.30)	74(0.23)	
≥18	98(0.73)	116(0.73)		264(0.77)	133(0.75)	124(0.89)		77(0.71)	74(0.70)	249(0.77)	
Age (average years)#	34.35	33.51		36.84	36.19	40.22		35.47	31.81	36.49	
	(31.46,37.24)	(30.82,36.20)		(34.91,38.77)	(33.51,38.87)	(37.58,42.85)		(31.79,39.16)	(28.63,34.99)	(34.55,38.43)	J
Vaccinated in previous year											
Yes	9(0.07)	29(0.18)		7(0.02)	21(0.12)	23(0.17)	*	16(0.15)	18(0.17)	68(0.21)	
No	126(0.93)	130(0.82)		337(0.98)	157(0.88)	116(0.83)		92(0.85)	88(0.83)	255(0.79)	

^{*}p < 0.05.

between the comparison groups in Table 2. Although mostly insignificant, a higher risk of infection, indicated by odds ratio greater than 1, was consistently observed among household children contacts from the un-vaccinated households. During the first year, a significantly higher risk of infection and clinical morbidity was observed among children contacts from un-vaccinated households than from complete vaccinated households (p < 0.05). No statistical significance for risk of infection and morbidity was observed among household adult contacts and odds ratios were not consistently greater than 1.

Interestingly, after adjusting for households with or without TIV vaccination, lower risk of infection was observed among household adult contacts from the rural residential area compared to those from the urban area (OR = 0.89; 95% CI: 0.82-0.97 for Year 2 and OR = 0.85; 95% CI: 0.75-0.96 for Year 3) (Table 6). Similar results were also observed among adult contacts for infection with clinical symptoms.

Discussion

Although vaccine recommendations have been gradually expanded to include all persons ≥6 months of age, the vaccination strategy specifically targeting schoolchildren to reduce community-wide or household transmission of seasonal influenza is still debatable [20-22]. Recent school-based and community cluster-randomized control trials suggest that the vaccination of schoolchildren can reduce influenza-related morbidity and mortality among non-immunized contacts within households [14] and in the community [10-13]. Consistent with previous reports, secondary infection risks within households were highest among young contacts [23-25]. In our current study, although no statistically significant reduction of serological-confirmed infection or clinical morbidity among the vaccinated household contacts was observed during three consecutive seasons, a consistent trend was observed among household children contacts from un-vaccinated households having a greater risk of infection

^{*}Numbers in brackets are 95 % confidence interval calculated by GEE method.

Table 3 Hemagglutination antibody titers of schoolchildren participant sera obtained before, after vaccination and post-season during three consecutive years

	1st Year			2nd Year			3rd Year		
	TIV(-)	TIV(+)	Р	TIV(-)	TIV(+)	Р	TIV(-)	TIV(+)	Р
	N = 59	N = 65		N = 172	N = 99		N = 67	N = 162	
Pre-vaccination									
HI-titer≧40									
H1N1	=	=		155(0.90)	91(0.92)		45(0.67)	124(0.77)	
H3N2	=	=		142(0.83)	77(0.78)		17(0.25)	58(0.36)	
В	-	=		44(0.26)	26(0.26)		49(0.73)	104(0.64)	*
GMT [#]									
H1N1	-	=		263.78(211.80,315.75)	412.83(246.23,579.42)	*	82.76(60.50,105.02)	89.38(75.14,103.62)	
H3N2	-	=		210.26(168.06,252.46)	245.91(166.90,324.92)		52.54(20.37,84.70)	85.40(57.74,113.06)	
В	-	=		25.76(21.17,30.34)	30.81(22.25,39.36)		69.18(53.71,84.65)	75.43(54.95,95.92)	
Post-vaccination									
HI-titer≧40									
H1N1	33(0.56)	60(0.92)	*	112(0.65)	61(0.62)		37(0.55)	132(0.81)	*
H3N2	51(0.86)	59(0.91)		114(0.66)	61(0.62)		6(0.09)	49(0.30)	*
В	26(0.44)	43(0.66)	*	72(0.42)	39(0.39)		23(0.34)	73(0.45)	
GMT#									
H1N1	161.80(91.12,232.48)	422.77(232.82,612.71)	*	408.08(306.32,509.84)	324.14(245.96,402.32)		97.76(19.45,176.07)	121.33(100.60,142.06)	
H3N2	353.93(253.05,454.81)	411.41(312.73,510.09)		487.59(407.64,567.53)	455.61(368.75,542.46)		20.82(12.91,28.73)	61.82(41.07,82.57)	*
В	115.74(67.21,164.28)	114.24(83.34,145.14)		46.48(38.09,54.87)	48.59(36.73,60.45)		33.13(24.06,42.21)	44.26(33.37,55.15)	
4-Fold increase (F	Post vs Pre-vaccination)								
H1N1	-	-		39(0.23)	17(0.17)		11(0.16)	82(0.51)	*
H3N2	-	-		60(0.35)	36(0.36)		3(0.04)	15(0.09)	
В	-	-		62(0.36)	31(0.31)		0(0.00)	26(0.16)	*
Post-season									
HI-titer≧40									
H1N1	40(0.68)	60(0.68)		154(0.90)	89(0.90)		45(0.67)	116(0.72)	
H3N2	48(0.82)	59(0.82)		163(0.95)	95(0.96)		28(0.42)	75(0.46)	
В	24(0.40)	43(0.40)		167(0.97)	93(0.94)		26(0.39)	68(0.42)	
GMT [#]	-	-							
H1N1	129.26(71.17,187.34)	157.46(68.66,246.25)		570.61(497.40,643.82)	503.08(410.42,595.74)		132.69(87.48,177.89)	90.52(77.27,103.78)	*
H3N2	272.46(159.43,385.48)	209.03(120.55,297.52)		319.04(269.30,368.79)	390.61(316.43,464.78)		99.63(34.88,164.38)	165.77(110.13,221.41)	
В	40.43(28.70,52.15)	55.53(38.39,72.67)		23.46(20.42,26.50)	24.70(20.09,29.31)		31.34(22.69,40.00)	37.87(28.82,46.92)	

^{*}p < 0.05.

^{*}Numbers in brackets are 95% confidence interval calculated by GEE method.

Table 4 Hemagglutination antibody titers during the second influenza season based on TIV or pH1N1 vaccination status of schoolchildren participants

Participants	TI	V(-) N = 172	TIV(+) N = 99					
(schoolchildren)	pH1N1(-) N = 17	pH1N1(+) N = 155	P	pH1N1(-) N = 6	pH1N1(+) N = 93	P		
Pre-vaccination								
HI-titer≧40								
pH1N1	11(0.65)	112(0.72)		4(0.67)	68(0.73)			
H1N1	14(0.82)	141(0.91)		4(0.67)	87(0.94)			
H3N2	13(0.77)	129(0.83)		5(0.83)	72(0.77)	*		
В	5(0.29)	39(0.25)		2(0.33)	24(0.26)			
GMT [#]								
pH1N1	48.24(28.72,67.75)	56.77(49.45,64.10)		65.00(7.23,122.77)	57.58(48.92,66.24)			
H1N1	162.94(79.96,245.92)	274.84(217.97,331.71)	*	433.33(-64.89,931.55)	411.51(235.43,587.58)			
H3N2	207.06(42.89,371.22)	210.61(166.67,254.56)		107.50(40.65,174.35)	254.84(171.02,338.65)	*		
В	17.65(9.77,25.52)	26.65(21.64,31.65)	*	66.67(-64.31,197.64)	28.49(21.90,35.09)			
Post-vaccination								
HI-titer <u>≥</u> 40								
pH1N1	2(0.12)	101(0.65)	*	2(0.33)	52(0.56)	*		
H1N1	4(0.24)	108(0.70)	*	3(0.50)	58(0.62)	*		
H3N2	4(0.24)	110(0.71)	*	4(0.67)	57(0.61)			
В	2(0.12)	70(0.45)	*	3(0.50)	36(0.39)			
GMT [#]								
H1N1_p	49.41(30.22,68.60)	216.65(167.73,265.56)	*	84.17(14.68,153.65)	228.60(154.87,302.33)	*		
H1N1	202.94(40.19,365.69)	430.58(319.28,541.88)	*	270.00(45.61,494.39)	327.63(245.00,410.27)			
H3N2	220.00(58.35,381.65)	516.94(430.86,603.01)	*	213.33(-9.85,436.51)	471.24(380.18,562.29)	*		
В	23.24(10.48,35.99)	49.03(39.89,58.17)	*	148.33(-3.62,300.29)	42.15(33.09,51.21)			

^{*}p < 0.05.

or combined clinical morbidity (odds ratio greater than 1) after adjusting for confounding variables. During Year 3, odds ratio less than 1 was observed among children contacts from partial-immunization households, which was probably due to the small sample size. In contrast, no consistent trend was observed among household adult contacts since the resulting odds ratios were not always greater than 1. One reason could be that household adults usually acquire infection from the community [15,26]. In particular, we observed a higher rate of infection in adults living in urban areas than those living in rural areas. Therefore, vaccinating schoolchildren may not provide sufficient protection to the household adult contacts living in urban cities. Since we didn't measure risk of infection conditioned on the level of exposure, only overall risk was considered in this study by adjusting the confounding variables using GEE methods and the potential bias of household vaccination status might affect the risk of exposure cannot be ruled out; ie. the household contact from unvaccinated households tended to be more careful to prevent having influenza infection. However, higher risk of infection among household adult contacts was observed in the urban area with lower vaccination rate, than rural area with higher vaccination rate (Tables 2 and 6). In the future, large clinical trials considering the transmission intensity in different geographic locations will be needed to assess the benefits of the schoolchildren vaccination policy on the household contacts and community.

Observational studies to assess vaccine effectiveness have become a standard way of routinely evaluating how well influenza vaccines protect population groups in countries implementing influenza vaccination programs [27-30]. These studies usually utilize a sensitive and specific laboratory method such as virus isolation or reverse-transcriptase polymerase chain reaction (RT-PCR) to confirm illness as influenza. These methods may miss a case if administered too early (e.g., during the incubation period) or too late (e.g., after the infectious period) in the disease course. Since many influenza cases are asymptomatic, identifying every case would necessitate serially screening the entire study population throughout the influenza season, which

^{*}numbers in brackets are 95% confidence interval calculated by GEE method.

Table 5 Proportions of infection or clinical morbidity among household contacts of children and adults among three TIV vaccination status groups during the three consecutive influenza seasons

	1st Year			2nd Year				3rd Year			
Children, Number (proportion)	TIV(-)	TIV(+)	Р	TIV(-)	Partial TIV	TIV(+)	Р	TIV(-)	Partial TIV	TIV(+)	Р
	N = 37	N = 43		N = 72	N = 38	N = 14		N = 31	N = 32	N = 74	
4-fold increase											
H1N1	6(0.16)	6(0.14)		29(0.40)	10(0.26)	3(0.21)		13(0.42)	12(0.38)	31(0.42)	
H3N2	4(0.11)	14(0.33)		30(0.42)	17(0.45)	4(0.29)		11(0.35)	6(0.19)	26(0.35)	
В	4(0.11)	3(0.07)		35(0.49)	17(0.45)	7(0.50)		6(0.19)	14(0.44)	18(0.24)	
pH1N1	-	-		55(0.76)	27(0.71)	13(0.93)		-	-	-	
Any 4-fold increase	14(0.38)	20(0.47)		44(0.61)	23(0.61)	7(0.50)		22(0.71)	21(0.66)	52(0.70)	
Symptoms											
Any	17(0.46)	29(0.67)		36(0.50)	22(0.58)	9(0.64)		19(0.61)	13(0.41)	37(0.50)	
Fever	5(0.14)	14(0.33)		7(0.10)	1(0.03)	1(0.07)		3(0.10)	6(0.19)	9(0.12)	
Fever plus any two symptoms	3(0.08)	10(0.23)		4(0.06)	0(0.00)	0(0.00)		1(0.03)	4(0.13)	7(0.09)	
Any 4-fold increase plus any symptoms	5(0.14)	15(0.35)	*	24(0.33)	13(0.34)	5(0.36)		27(0.87)	24(0.75)	62(0.84)	
Adults, Number (proportion)	TIV(-)	TIV(+)	Р	TIV(-)	Partial TIV	TIV(+)	Р	TIV(-)	Partial TIV	TIV(+)	Р
	N = 98	N = 116		N = 222	N = 111	N = 109		N = 77	N = 74	N = 249	
Any 4-fold increase											
H1N1	0(0.00)	2(0.02)		70(0.32)	46(0.41)	41(0.38)		10(0.13)	19(0.26)	60(0.24)	
H3N2	9(0.09)	17(0.15)		57(0.26)	38(0.34)	41(0.38)	*	3(0.04)	0(0.00)	6(0.02)	
В	5(0.05)	3(0.03)		70(0.32)	45(0.41)	38(0.35)		4(0.05)	6(0.08)	18(0.07)	
pH1N1	-	-		109(0.49)	57(0.51)	55(0.50)		-	-	-	
Any 4-fold increase	14(0.14)	20(0.17)		135(0.61)	74(0.67)	75(0.69)		11(0.14)	20(0.27)	64(0.26)	
Symptoms											
Any	43(0.44)	52(0.45)		55(0.25)	34(0.31)	39(0.36)	*	30(0.39)	31(0.42)	105(0.42)	
Fever	9(0.09)	5(0.04)		1(0.00)	4(0.04)	2(0.02)		2(0.03)	1(0.01)	7(0.03)	
Fever plus any two symptoms	0(0.00)	1(0.01)		0(0.00)	0(0.00)	0(0.00)		1(0.01)	0(0.00)	5(0.02)	
Any 4-fold increase plus any symptoms	7(0.07)	7(0.06)		39(0.18)	24(0.22)	24(0.22)		38(0.49)	42(0.57)	131(0.53)	

^{*}P < 0.05.

currently would be prohibitively expensive and operationally challenging. The advantage of using serological four-fold increase of HI titers in this study as a surrogate marker of infection was that the true infection rate could be estimated. To measure the true infections among household contacts, all contacts having TIV vaccination records during the contemporary influenza season were pre-excluded from the analysis. The possibility of a delayed rising of antibody titer post-infection, resulting in mis-classification of infection status among different groups, was also considered. The median durations of sera withdrawn were equivalent among household contacts from households with different vaccination statuses. As shown in Table 5, the infection rates from serological results among the household contacts were much higher than those previously published in studies using RT-PCR to estimate infection [12,22,31]; however, consistent with our previous publication [32]. All virus strains used in this study were originated from human, instead of avianorigin, and the negative control was used all the time during performing the assay. Although guinea pig RBC is strongly suggested for HI assay, the potential bias using chicken RBC in this study will be minimal but cannot be completely excluded [33,34]. However, the disadvantage of using serological four-fold increase of HI titers as an infection marker was that if the true benefits of the schoolchildren vaccinating policy is to prevent clinical illness instead of infection, this method under-estimated the vaccination effectiveness. Therefore, no statistically significant reduction of serological-confirmed infection among the vaccinated household contacts was observed during three consecutive seasons in this study due to the high infection rate by influenza virus. The unexpected findings of high infection rate and low effectiveness of household contacts require further study.

Measuring serologically confirmed infection allowed us to compare the transmission and infection of influenza virus in different geographical locations. Our multivariate

Table 6 Multivariate logistic regression on proportion of infection or plus clinical morbidity among household contacts during three consecutive influenza seasons

		Any 4-fold increase ^{\$}	р	Any 4-fold increase plus any symptoms ^{\$}	р
Children					
1st Year	Rural v.s urban	1.03 (0.79,1.34)	0.835	0.93 (0.76,1.13)	0.458
	Household size	0.97 (0.85,1.11)	0.692	0.96 (0.87,1.06)	0.385
	TIV previous year	0.72 (0.42,1.22)	0.221	0.97 (0.76,1.25)	0.831
	TIV (-) v.s (+)	1.26 (0.93,1.70)	0.139	1.27 (1.03,1.56)	0.024*
2nd Year	Rural v.s urban	0.95 (0.79,1.13)	0.547	0.93 (0.76,1.14)	0.505
	Household size	1.01 (0.81,1.27)	0.901	1.00 (0.79,1.27)	0.997
	TIV previous year	0.91 (0.60,1.38)	0.664	1.10 (0.70,1.75)	0.672
	TIV (-) v.s (+)	1.08 (0.76,1.52)	0.664	1.03 (0.70,1.52)	0.886
	TIV partial v.s (+)	1.09 (0.75,1.58)	0.665	1.06 (0.68,1.66)	0.800
3rd Year	Rural v.s urban	1.10 (0.96,1.25)	0.171	1.02 (0.84,1.24)	0.834
	Household size	1.03 (0.93,1.13)	0.607	1.07 (0.95,1.21)	0.260
	TIV this year	1.03 (0.91,1.16)	0.665	0.94 (0.79,1.11)	0.451
	TIV (-) v.s (+)	1.09 (0.96,1.24)	0.184	1.13 (0.92,1.38)	0.249
	TIV partial v.s (+)	0.91 (0.76,1.09)	0.306	0.85 (0.66,1.08)	0.186
Adults					
1st Year	Rural v.s urban	0.98 (0.87,1.21)	0.161	1.02 (0.93,1.12)	0.693
	Household size	1.02 (0.97,1.07)	0.444	1.00 (0.97,1.04)	0.806
	TIV previous year	1.11 (0.91,1.35)	0.314	0.94 (0.90,0.97)	<0.001*
	TIV (-) v.s (+)	1.01 (0.91,1.12)	0.878	0.99 (0.92,1.06)	0.726
2nd Year	Rural v.s urban	0.89 (0.82,0.97)	0.006*	0.91 (0.84,0.99)	0.024*
	Household size	0.89 (0.80,1.00)	0.046*	0.99 (0.90,1.08)	0.762
	TIV previous year	0.98 (0.83,1.16)	0.809	1.08 (0.92,1.28)	0.342
	TIV (-) v.s (+)	0.95 (0.85,1.06)	0.359	0.98 (0.89,1.08)	0.688
	TIV partial v.s (+)	1.07 (0.94,1.23)	0.303	1.02 (0.91,1.16)	0.700
3rd Year	Rural v.s urban	0.85 (0.75,0.96)	0.010*	0.87 (0.77,0.99)	0.031*
	Household size	1.03 (0.93,1.13)	0.620	1.05 (0.94,1.18)	0.385
	TIV previous year	1.08 (0.91,1.30)	0.381	0.96 (0.80,1.14)	0.627
	TIV (-) v.s (+)	0.94 (0.81,1.09)	0.423	0.95 (0.82,1.10)	0.476
	TIV partial v.s (+)	0.99 (0.83,1.19)	0.919	0.92 (0.75,1.12)	0.402

^{\$}Odds ratio (95% confidence interval).

GEE analysis after adjusting household vaccination status suggested that a lower risk of acquiring infection was observed among household adult contacts from the rural residential area than among those from the urban area with statistical significance (Table 6). Recent studies from 2009 pandemic H1N1 (pH1N1) outbreaks suggested that the transmissions of influenza virus were spatially heterogeneous [15,26]. A different contact profile in different social settings, social and geographic factors probably shape the local reproduction number (R_0) [35,36]. The crowdedness of urban environments facilitates influenza virus transmission, and adults might acquire infection through variable contact pathways in the community. As such, this

might compromise the vaccine effectiveness of the schoolchildren vaccination policy in reducing influenza-related morbidity and mortality of household contacts, particularly in adults. Designing an optimal mitigation measurement tailored to different geographical locations will help to control the influenza epidemic or pandemic.

Another interesting finding in our study was that higher antibody titers against seasonal influenza strains including H1, H3 and B were observed in children who received the pH1N1 vaccine than in those who did not. This observation was independent from whether or not the children received TIV. Previous studies suggested that broadly neutralizing antibodies reactive to the

^{*}p < 0.05.

conserved stem region of the influenza virus hemagglutinin (HA) were generated in people infected or vaccinated with the 2009 pandemic H1N1 strain [37,38]. Additionally, other studies also suggested that a vigorous antibody response to pH1N1 vaccination resulted from the activation of preexisting memory B cells, which did not only bind to the stem region but also the head region of HA [39]. Furthermore, cross-reactive T-cell responses can also contribute to the cross-protection. A larger sample size will be needed to confirm this observation.

The main limitation of our study relates to the small sample sizes among subsets of the participants receiving TIV, which reduced both the statistical power and precision. Same reason for the attempt of calculating attack rates within households and compare them among vaccinated, partial and non-vaccinated households. Particularly in Year 2, two vaccines (TIV and pH1N1) were delivered during the same season one month apart. Our previous study demonstrated that children who received TIV prior to pH1N1 had a lower sero-conversion rate than those who received only the pH1N1 vaccine [40]. Whether this was the cause of the higher infection rates among the household contacts observed in our study, as shown in Table 5, requires further confirmation. Potential self-selection bias of study subjects involved in this study is possible, particularly since more families were recruited from rural areas than urban areas in this study. Furthermore, timing of vaccination or period of sera collection coincided with the circulation of influenza viruses might also affect the determination of status of infection, particularly during Year 2. Prior immunity of each individual could also affect the protection level, data of which was available for this study but appeared insignificant due to small sample size after stratification.

Conclusions

Although drawing a conclusion from data in this observational study requires careful interpretation of the results, due to the possibility of bias of given self-selection for vaccination, assessing the effectiveness of the school-children vaccination policy on protection of household contacts and community would require further research on geographical heterogeneity of influenza transmission and infection.

Competing interests

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Authors' contributions

DYC designed the laboratory work and decided the virus strains carried out in this study. CYC and TCL participated in the design of the study, designed the questionnaire and coordinated the training for telephone interview. DYC, KFC, YHH, and TNW conceived of the study, statistical analysis and

participated in its design and coordination and helped to draft the manuscript. All authors read and approved the final manuscript.

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Author details

¹Graduate Institute of Microbiology and Public Health, College of Veterinary Medicine, National Chung-Hsing University, Taichung, Taiwan. ²School of Public Health and Biostatistics Center, Taipei Medical University, Taipei, Taiwan. ³Graduate Institute of Biostatistics, China Medical University, Taichung, Taiwan. ⁴Department of Public Health, China Medical University, Taichung, Taiwan.

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